
Gaussian Measures in Euclidean Space

1.1. The change of variables formula

Let (X, S, μ) be a measure space, i.e. X is a non-empty set, S is a sigma-algebra on X and μ is a measure on S . Consider also a measurable space (Y, F) , i.e. Y is another non-empty set and F is a sigma-algebra on Y . Let $T : X \rightarrow Y$ be a measurable transformation, which means that

$$\forall A \in F, \quad T^{-1}A \in S. \quad [1.1]$$

Hereafter

$$T^{-1}A := \{x \in X : Tx \in A\} \quad [1.2]$$

is preimage of A under T . (To simplify the notation, we write $T^{-1}A$ and Tx rather than $T^{-1}(A)$ and $T(x)$, respectively, if it does not cause confusion.)

Introduce a set function

$$\nu(A) = \mu(T^{-1}A), \quad A \in F. \quad [1.3]$$

THEOREM 1.1.– (About induced measure) The set of function ν given in [1.3] is a measure on F .

PROOF.– The function ν is well defined due to [1.1]. We have to show that it is not identical to infinity, but it is non-negative and sigma-additive.

Indeed, $\nu(\emptyset) = \mu(T^{-1}\emptyset) = \mu(\emptyset) = 0$, and therefore, ν is not identical to infinity. For each $A \in F$, $\nu(A) \geq 0$, because μ is a non-negative set function.

Finally, $\{A_n, n \geq 1\}$ are disjoint sets from F . Then the preimages $\{T^{-1}A_n, n \geq 1\}$ are disjoint as well, and

$$\begin{aligned} \nu \left(\bigcup_{n=1}^{\infty} A_n \right) &= \mu \left(T^{-1} \bigcup_{n=1}^{\infty} A_n \right) = \mu \left(\bigcup_{n=1}^{\infty} T^{-1}A_n \right) = \\ &= \sum_{n=1}^{\infty} \mu (T^{-1}A_n) = \sum_{n=1}^{\infty} \nu(A_n). \end{aligned}$$

Here, we used the sigma-additivity of μ . Thus, ν is a non-negative and sigma-additive set function on the sigma-algebra F , i.e. ν is a measure on F . \square

DEFINITION 1.1.– *The set function ν given in [1.3] is called a measure induced by transformation T and is denoted as μT^{-1} .*

The notation prompts how to evaluate $\nu(A)$:

$$(\mu T^{-1})(A) = \mu(T^{-1}A), \quad A \in F.$$

For any measurable space (Y, F, ν) , denote $L(F, \nu)$ the space of Lebesgue integrable functions on Y w.r.t. measure ν .

Let $f : Y \rightarrow \bar{\mathbb{R}}$ be an F -measurable function, i.e. for each Borel subset B of extended real line $\bar{\mathbb{R}}$, it holds $f^{-1}B \in F$.

THEOREM 1.2.– (The change of variables formula) Assume that either $f \geq 0$ or $f \in L(Y, \mu T^{-1})$. Then it holds

$$\int_X f(Tx) d\mu(x) = \int_Y f(y) d(\mu T^{-1})(y). \tag{1.4}$$

PROOF.– Equality [1.4] is shown in a standard way: first for indicators, then for simple non-negative functions, then for $f \geq 0$, and finally, for $f \in L(Y, \mu T^{-1})$.

a) Let $A \in F$,

$$f(y) = I_A(y) = \begin{cases} 1, & \text{if } y \in A \\ 0, & \text{otherwise.} \end{cases}$$

Then

$$I_A(Tx) = \begin{cases} 1, & \text{if } Tx \in A \\ 0, & \text{otherwise,} \end{cases}$$

$$I_A(Tx) = \begin{cases} 1, & \text{if } x \in T^{-1}A \\ 0, & \text{otherwise,} \end{cases}$$

$$I_A(Tx) = I_{T^{-1}A}(x).$$

Hence

$$\begin{aligned} \int_X I_A(Tx) d\mu(x) &= \int_X I_{T^{-1}A}(x) d\mu(x) = \mu(T^{-1}A) = (\mu T^{-1})(A), \\ \int_Y I_A(y) d(\mu T^{-1})(y) &= (\mu T^{-1})(A), \end{aligned}$$

and [1.4] follows for the indicator function.

b) Let $f \geq 0$ be a simple F -measurable function. Then it admits a representation

$$f(y) = \sum_{k=1}^m a_k I_{A_k}(y), \quad y \in Y, \quad [1.5]$$

with disjoint measurable sets $\{A_k, k = 1, \dots, m\}$ and non-negative a_k . For the function [1.5], relation [1.4] follows due to part (a) of the proof and the linearity of the Lebesgue integral.

c) Let f be an arbitrary non-negative and F -measurable function. Then there exists a sequence $\{p_n(y), n \geq 1, y \in Y\}$ of non-negative, simple and F -measurable functions such that p_n converges to f pointwise and $p_n(y) \leq p_{n+1}(y), n \geq 1, y \in Y$. By part (b) of the proof,

$$\int_X p_n(Tx) d\mu(x) = \int_Y p_n(y) d(\mu T^{-1})(y). \quad [1.6]$$

Here, tend n to infinity. By the monotone convergence theorem, [1.6] implies [1.4].

d) Finally, let $f \in L(Y, \mu T^{-1})$,

$$f_+(y) := \max\{f(y), 0\}, \quad f_-(y) := -\min\{f(y), 0\}, \quad y \in Y.$$

By part (c) of the proof,

$$\int_X f_+(Tx) d\mu(x) = \int_Y f_+(y) d(\mu T^{-1})(y), \quad [1.7]$$

$$\int_X f_-(Tx) d\mu(x) = \int_Y f_-(y) d(\mu T^{-1})(y). \quad [1.8]$$

Subtract [1.8] from [1.7] and obtain [1.4] using the definition of Lebesgue integral.

□

Problems 1.1

1) Let λ_2^T be a Lebesgue measure on $[0, T]^2$; $\pi_1(x_1, x_2) = x_1$, $(x_1, x_2) \in [0, T]^2$, $\pi_1 : [0, T]^2 \rightarrow \mathbb{R}$. Show that

$$(\lambda_2^T \pi_1^{-1})(A) = T \cdot \lambda_1(A \cap [0, T]), \quad A \in S_1,$$

where λ_1 is Lebesgue measure on \mathbb{R} and S_1 is sigma-algebra of Lebesgue measurable sets on \mathbb{R} .

2) Let μ_1 and μ_2 be finite measures on Borel sigma-algebra $\mathcal{B}(\mathbb{R})$, and $\pi(x_1, x_2) = x_1$, $(x_1, x_2) \in \mathbb{R}^2$. Find the induced measure $(\mu_1 \times \mu_2)\pi_1^{-1}$.

3) For the objects of theorem 1.1, prove the following: if μT^{-1} is sigma-finite, then μ is sigma-finite as well. Does the opposite hold true?

4) Let $f : Y \rightarrow \bar{\mathbb{R}}$ be any F -measurable function. Show that the Lebesgue integral on the left-hand side of [1.4] is well defined if, and only if, the integral on the right-hand side of [1.4] is well defined. Moreover, in case where they are well defined, they coincide.

1.2. Invariance of Lebesgue measure

Consider a measure space (X, S, μ) and a measurable transformation $T : X \rightarrow X$.

DEFINITION 1.2.— *The measure μ is called invariant under T , or T -invariant, if $\mu T^{-1} = \mu$.*

REMARK 1.1.— Assume additionally that T is a bijection on X , and moreover T^{-1} is a measurable transformation as well. Then μ is T -invariant if and only if

$$\mu(B) = \mu(TB), \quad \forall B \in S. \quad [1.9]$$

(Hereafter TB denotes image of B under T .)

PROOF.—

a) Let μ be T -invariant and $B \in S$. Because T^{-1} is measurable, $A := TB \in S$. It holds $B = T^{-1}A$, and $\mu(T^{-1}A) = \mu(A)$. Equality [1.9] follows.

b) Conversely, assume [1.9] and take any $A \in S$. Denote $B_0 = T^{-1}A$, $B_0 \in S$. Then $(\mu T^{-1})(A) = \mu(B_0) = \mu(TB_0) = \mu(A)$, and μ is T -invariant. \square

EXAMPLE 1.1.— (Counting measure) Let $X = \{1, 2, \dots, n\}$, $S = 2^X$ be the sigma-algebra of all subsets of X , and μ be the counting measure on X , i.e. $\mu(A) = |A|$, $A \in S$. (Hereafter $|A|$ is number of points in a set A ; if A is infinite, $|A| = +\infty$.) Then

μ is invariant under any bijection π on X . Indeed, $\mu(\pi^{-1}A) = |\pi^{-1}A| = |A| = \mu(A)$, $A \in S$.

In this section, we show that Lebesgue measure λ_n on \mathbb{R}^n is rotation and translation invariant. Hereafter, we suppose that Euclidean space \mathbb{R}^n consists of column vectors $x = (x_1, \dots, x_n)^\top$.

DEFINITION 1.3.— *Affine transformation of \mathbb{R}^n is a mapping of a form $Tx = Lx + c$, with $L \in \mathbb{R}^{n \times n}$ and $c \in \mathbb{R}^n$. Such transformation is called non-singular if L is non-singular. Otherwise, if $\det L = 0$, then T of this form is called a singular affine transformation.*

REMARK 1.2.— Affine transformation $Tx = Lx + c$ is invertible if, and only if, it is non-singular. In this case, the inverse transformation is a non-singular affine transformation as well, and it acts as follows:

$$T^{-1}y = L^{-1}y - L^{-1}c, \quad y \in \mathbb{R}^n.$$

Remember that non-singular affine transformations on a plane include rotations, translations and axial symmetries.

THEOREM 1.3.— (Transformation of Lebesgue measure at Borel sets) Consider Lebesgue measure λ_n on Borel sigma-algebra $\mathcal{B}(\mathbb{R}^n)$. Let $Tx = Lx + c$ be a non-singular affine transformation on \mathbb{R}^n . Then

$$\lambda_n T^{-1} = \frac{1}{|\det L|} \lambda_n. \quad [1.10]$$

PROOF.— Transformation T is continuous, and, therefore, Borel measurable. Hence the induced measure $\lambda_n T^{-1}$ on $\mathcal{B}(\mathbb{R}^n)$ is well defined.

a) For $a = (a_k)_1^n$ and $b = (b_k)_1^n$ with $a_k < b_k$, $k = 1, \dots, n$ denote

$$[a, b] = \prod_{k=1}^n [a_k, b_k], \quad (a, b) = \prod_{k=1}^n (a_k, b_k). \quad [1.11]$$

Hereafter, $\prod_{k=1}^n A_k$ stands for Cartesian product of A_1, \dots, A_n . Evaluate

$$(\lambda_n T^{-1})([a, b]) = \lambda_n(T^{-1}[a, b]) = \int_{T^{-1}[a, b]} d\lambda_n = \int_{T^{-1}[a, b]} dx.$$

The latter integral is Riemann integral over the compact and Jordan measurable set $T^{-1}[a, b]$. The change of variables in the Riemann integral leads to the following:

$$\lambda_n(T^{-1}[a, b]) = \int_{[a, b]} \left| \frac{\partial y}{\partial x} \right|^{-1} dy = \frac{m_n([a, b])}{|\det L|} = \frac{\lambda_n([a, b])}{|\det L|}.$$

Here m_n is Jordan measure on \mathbb{R}^n .

b) Consider a set $(a, b]$ introduced in [1.11], and let $\{a_k(m), m \geq 1\}$ be a decreasing sequence that converges to a_k such that $a_k(m) < b_k, m \geq 1; k = 1, \dots, n$. Denote $a(m) = (a_k(m))_{k=1}^n \in \mathbb{R}^n$. Then $A_m := [a(m), b]$ is a monotone sequence of sets that converges to $(a, b]$. The continuity of Lebesgue measure from below implies

$$(\lambda_n T^{-1})((a, b]) = \lim_{m \rightarrow \infty} \lambda_n(T^{-1}A_m) = \lim_{m \rightarrow \infty} \frac{\lambda_n(A_m)}{|\det L|} = \frac{\lambda_n((a, b])}{|\det L|}.$$

Here, we used part (a) of the proof.

c) Thus, the two measures $\lambda_n T^{-1}$ and $\frac{\lambda_n}{|\det L|}$ in [1.10] coincide on the semiring \mathcal{P}_n that consists of all bricks $(a, b]$ from [1.11]. Both measures are sigma-finite, and therefore, they coincide on $\sigma r(\mathcal{P}_n) = \mathcal{B}(\mathbb{R}^n)$, where $\sigma r(\mathcal{P}_n)$ is sigma-ring generated by \mathcal{P}_n . \square

Now, we extend theorem 1.3 to Lebesgue measure λ_n on sigma-algebra S_n of Lebesgue measurable sets in \mathbb{R}^n .

LEMMA 1.1.– Non-singular affine transformation $Tx = Lx + c$ is (S_n, S_n) -measurable, i.e. for any $A \in S_n, T^{-1}A \in S_n$ as well.

PROOF.– It is known (see [HAL 13]) that

$$S_n = \{B \cup N \mid B \in \mathcal{B}(\mathbb{R}^n), N \subset N_0 \text{ with } N_0 \in \mathcal{B}(\mathbb{R}^n), \lambda_n(N_0) = 0\}. \quad [1.12]$$

Let $A \in S_n$, then $A = B \cup N$, with B and N described in [1.12]. It holds

$$T^{-1}A = (T^{-1}B) \cup (T^{-1}N), \quad T^{-1}B \in \mathcal{B}(\mathbb{R}^n), \quad [1.13]$$

$$T^{-1}N \subset T^{-1}N_0, T^{-1}N_0 \in \mathcal{B}(\mathbb{R}^n), \lambda_n(T^{-1}N_0) = \frac{\lambda_n(N_0)}{|\det L|} = 0. \quad [1.14]$$

Here, we used theorem 1.3 and the fact that T is a Borel function. Decompositions [1.13] and [1.14] show that $T^{-1}A \in S_n$. \square

THEOREM 1.4.– (Transformation of Lebesgue measure) Let $Tx = Lx + c$ be a non-singular affine transformation on \mathbb{R}^n . For Lebesgue measure λ_n on S_n , it holds

$$\lambda_n T^{-1} = \frac{\lambda_n}{|\det L|}.$$

PROOF.– Consider $A \in S_n$ and decompose $T^{-1}A$ as in [1.13] and [1.14]. Because $\lambda_n(N) = \lambda_n(T^{-1}N) = 0$, we have by theorem 1.3:

$$\lambda_n(T^{-1}A) = \lambda_n(T^{-1}B) = \frac{\lambda_n(B)}{|\det L|} = \frac{\lambda_n(A)}{|\det L|}. \quad \square$$

COROLLARY 1.1.– (Criterion for invariance of Lebesgue measure) Lebesgue measure λ_n on S_n is invariant under a non-singular affine transformation $Tx = Lx + c$ if, and only if, $\det L = \pm 1$.

In particular, λ_n is symmetric around the origin, and it is invariant under translations $Tx = x + c$ and orthogonal transformations $Tx = Ux$, where U is an orthogonal matrix (i.e. $U^{-1} = U^T$), e.g. under symmetries w.r.t. hyperplanes that pass through the origin. In the planar case ($n = 2$), λ_2 is invariant under the transformation

$$Tx = \begin{pmatrix} 2x_1 \\ \frac{1}{2}x_2 \end{pmatrix}, \quad x \in \mathbb{R}^2.$$

Here, T is dilation along x_1 -axis with coefficient 2 and contraction along x_2 -axis with the same coefficient.

COROLLARY 1.2.– (Affine change of variables) Let $Tx = Lx + c$ be a non-singular affine transformation on \mathbb{R}^n and $f : \mathbb{R}^n \rightarrow \bar{\mathbb{R}}$ be a Lebesgue measurable function, which is either non-negative or belongs to $L(\mathbb{R}^n, \lambda_n)$. Then it holds

$$\int_{\mathbb{R}^n} f(Tx) d\lambda_n(x) = \frac{1}{|\det L|} \int_{\mathbb{R}^n} f(y) d\lambda_n(y).$$

PROOF.– Apply theorems 1.2 and 1.4:

$$\int_{\mathbb{R}^n} f(Tx) d\lambda_n(x) = \int_{\mathbb{R}^n} f(y) d(\lambda_n T^{-1})(y) = \frac{1}{|\det L|} \int_{\mathbb{R}^n} f(y) d\lambda_n(y). \quad \square$$

Problems 1.2

5) Let $Tx = |x|$, $x \in \mathbb{R}$. Find $\lambda_1 T^{-1}$.

6) Let $f : \mathbb{R} \rightarrow \mathbb{R}$ be a Lebesgue measurable function such that $\forall x, y \in \mathbb{R}$, $|f(x) - f(y)| \geq |x - y|$. Prove that f is (S_1, S_1) -measurable, where S_1 is sigma-algebra of Lebesgue measurable sets on real line.

7) Show that $\arctan x$ is an (S_1, S_1) -measurable function. Let $f : (-\frac{\pi}{2}, \frac{\pi}{2}) \rightarrow [0, +\infty]$ be a Lebesgue measurable function. Prove that

$$\int_{\mathbb{R}} f(\arctan x) d\lambda_1(x) = \int_{(-\frac{\pi}{2}, \frac{\pi}{2})} \frac{f(t)}{\cos^2 t} d\lambda_1(t).$$

8) Show that e^x is (S_1, S_1) -measurable function. Let $f : (0, \infty) \rightarrow [0, +\infty]$ be a Lebesgue measurable function. Prove that

$$\int_{\mathbb{R}} f(e^x) d\lambda_1(x) = \int_{(0, +\infty)} \frac{f(t)}{t} d\lambda_1(x).$$

9) Prove that $f(x) = \|x\|$, $x \in \mathbb{R}^n$ is (S_n, S_1) -measurable function.

10) Let $f : [0, +\infty) \rightarrow [0, +\infty]$ be a Lebesgue measurable function. Prove that the measure

$$\mu(A) = \int_A f(\|x\|) d\lambda_n(x), \quad A \in S_n$$

is invariant under unitary operators in \mathbb{R}^n .

11) Let μ be a measure on S_n , which is finite at each bounded set from S_n , absolutely continuous w.r.t. λ_n and invariant under unitary operators in \mathbb{R}^n . Prove that there exists a Borel function $f : [0, +\infty) \rightarrow [0, +\infty)$ such that representation from Problem 10 holds true.

Hint. Given a locally Lebesgue integrable function g on \mathbb{R}^n , a point $x \in \mathbb{R}^n$ is Lebesgue point if

$$\lim_{r \rightarrow 0^+} \frac{1}{\lambda_n(B(x, r))} \int_{B(x, r)} |g(y) - g(x)| d\lambda_n(y) = 0.$$

Hereafter, $B(x, r)$ is an open ball centered at x with radius r . Use the Lebesgue differentiation theorem [BOG 07] which states that, given any locally Lebesgue integrable function g on \mathbb{R}^n , almost every x is a Lebesgue point of g .

12) Let $g : \mathbb{R}^n \rightarrow \mathbb{R}$ be a Lebesgue measurable function such that $g(Tx) = g(x) \pmod{\lambda_n}$, for all unitary operators T in \mathbb{R}^n . Prove that there exists a Borel function $f : [0, +\infty) \rightarrow \mathbb{R}$, with $g(x) = f(\|x\|) \pmod{\lambda_n}$.

13) Let $\alpha > 0$ and $f \in L(\mathbb{R}, \lambda_1)$. Prove that $f(n^{1+\alpha}x) \rightarrow 0$ as $n \rightarrow \infty$ for almost all $x \in \mathbb{R}$. Extend this statement to functions from $L(\mathbb{R}^m, \lambda_m)$.

14) Let $f : \mathbb{R}^n \rightarrow \mathbb{R}$, $f \in L([0, +\infty), \lambda_1)$. Prove the following:

a) If f is an even function, then

$$\int_{\mathbb{R}} f d\lambda_1 = 2 \int_{[0, +\infty)} f d\lambda_1.$$

b) If f is an odd function, then $\int_{\mathbb{R}} f d\lambda_1 = 0$.

15) Let $f : [-1, 1] \rightarrow (0, +\infty)$ be a Lebesgue measurable function. Find the integral $\int_{[-1, 1]} \frac{f(x)}{f(x)+f(-x)} d\lambda_1(x)$.

1.3. Absence of invariant measure in infinite-dimensional Hilbert space

Let H be a real Hilbert space, with Borel sigma-algebra $\mathcal{B}(H)$. In this section, we search for a measure λ on $\mathcal{B}(H)$ with the following properties:

- i) λ is positive at each non-empty open set;
- ii) λ is finite at each bounded Borel set;
- iii) λ is invariant under each translation $Tx = x + c$, $x \in H$, with $c \in H$.

Remember that a linear operator U in H is called unitary if $\|Ux\| = \|x\|$, $x \in H$, and $R(U) = H$. An operator $U \in L(H)$ is unitary if, and only if, $U^* = U^{-1}$.

- iv) λ is invariant under each unitary operator in H .

Note that Lebesgue measure λ_n on $\mathcal{B}(\mathbb{R}^n)$ possesses the properties (i)–(iv).

THEOREM 1.5.– (Absence of invariant measure in H) Let H be an infinite-dimensional real Hilbert space. Then:

- a) There is no measure λ with properties (i)–(iii).
- b) There is no measure λ with properties (i), (ii) and (iv).

PROOF.– Because $\dim(H) = \infty$, there exists an infinite orthonormal system $\{e_n, n \geq 1\}$ in H . For $k \neq m$, $\|e_k - e_m\| = \sqrt{2}$, hence open balls $B\left(e_n, \frac{\sqrt{2}}{2}\right)$ are disjoint. For $x \in B\left(e_n, \frac{\sqrt{2}}{2}\right)$, it holds $\|x\| \leq \|e_n\| + \|x - e_n\| < 1 + \frac{\sqrt{2}}{2} < 2$, and

$$B\left(e_n, \frac{\sqrt{2}}{2}\right) \subset B(0, 2), \quad \bigcup_{n=1}^{\infty} B\left(e_n, \frac{\sqrt{2}}{2}\right) \subset B(0, 2). \quad [1.15]$$

a) Let λ have the properties (i)–(iii). For $k \neq m$, the translation $Tx = x + e_m - e_k$, $x \in H$ maps the ball $B\left(e_k, \frac{\sqrt{2}}{2}\right)$ onto $B\left(e_m, \frac{\sqrt{2}}{2}\right)$. Hence by (i) and (iii),

$$\lambda\left(B\left(e_k, \frac{\sqrt{2}}{2}\right)\right) = \lambda\left(B\left(e_m, \frac{\sqrt{2}}{2}\right)\right) = a > 0. \quad [1.16]$$

Due to [1.15] and [1.16], we have

$$\lambda(B(0, 2)) \geq \sum_{n=1}^{\infty} \lambda\left(B\left(e_n, \frac{\sqrt{2}}{2}\right)\right) = \sum_{n=1}^{\infty} a = +\infty, \quad \lambda(B(0, 2)) = +\infty.$$

But this contradicts property (ii). Therefore, such a measure λ does not exist.

b) Now, assume that λ has the properties (i), (ii) and (iv). Construct a unitary operator U that maps $B\left(e_k, \frac{\sqrt{2}}{2}\right)$ onto $B\left(e_m, \frac{\sqrt{2}}{2}\right)$, with fixed $k \neq m$.

Let L be a subspace generated by $\{e_n, n = 1, 2, \dots\}$. Each $x \in H$ can be decomposed as

$$x = \sum_{n=1}^{\infty} (x, e_n) e_n + z, \quad z \in L^\perp.$$

The isometric operator

$$Ux = (x, e_k) e_m + (x, e_m) e_k + \sum_{n \neq k, n \neq m} (x, e_n) e_n + z$$

is a surjection, and hence is unitary. It maps $B\left(e_k, \frac{\sqrt{2}}{2}\right)$ onto $B\left(e_m, \frac{\sqrt{2}}{2}\right)$, and by properties (iv) and (i), it holds [1.16]. The rest of the proof follows the line of part (a) of the proof. \square

As we see, in space l_2 of sequences and in the space $L_2[a, b]$ of functions there is no measure analogous to Lebesgue measure. Nevertheless, we will construct a measure in an infinite-dimensional Hilbert space, which is invariant under quite a large group of transformations. It will be a Gaussian measure.

Problems 1.3

16) Prove that there is no measure λ on $\mathcal{B}(l_\infty)$ with properties (i) and (ii) from section 1.3, where l_∞ is the space of real bounded sequences with the supremum norm.

17) Let X be a real normed space, with $\dim(X) = \infty$. Prove that there is no measure λ on $\mathcal{B}(X)$ with properties (i)–(iii) from section 1.3.

18) A linear bijection V on a normed space X is called isometry if $\|Vx\| = \|x\|$, $x \in X$. Prove that there is no measure λ on $\mathcal{B}(l_p)$ with properties (i) and (ii) from section 1.3 and such that λ is invariant under all isometries on l_p , $1 \leq p < \infty$.

19) Let $\varphi(t)$, $t \in [0, 1]$ be a continuous increasing function, with $\varphi(0) = 0$, $\varphi(1) = 1$, and $\varphi(t) < t$, $t \in (0, 1)$. In Banach space, $X = C[0, 1]$ introduce a transformation $(Tx)(t) = x(\varphi(t))$, $t \in [0, 1]$, $x \in X$. Prove that there is no measure λ on $\mathcal{B}(X)$ with properties (i) and (ii) from section 1.3 and such that it is T -invariant.

1.4. Random vectors and their distributions

Remember that a probability measure is a measure on sigma-algebra, which is equal to 1 at the total space. A measure space (Ω, \mathcal{F}, P) is called probability space if P is a probability measure, i.e. $P(\Omega) = 1$.

1.4.1. Random variables

A random variable (r.v.) on a probability space (Ω, \mathcal{F}, P) is just an \mathcal{F} -measurable function on Ω .

DEFINITION 1.4.– Let $X = X(\omega)$ be a r.v. on a probability space (Ω, \mathcal{F}, P) . The distribution of X is a probability measure μ_X defined as follows:

$$\mu_X(B) = P\{\omega : X(\omega) \in B\}, \quad B \in \mathcal{B}(\mathbb{R}).$$

Note that μ_X is a measure induced by the mapping $X : \Omega \rightarrow \mathbb{R}$, i.e. $\mu_X = P X^{-1}$ (see definition 1.1).

A Borel function $f : \mathbb{R} \rightarrow \mathbb{R}$ is called the probability density function (pdf) of a r.v. X if

$$P\{X(\omega) \in B\} = \int_B f(t) d\lambda_1(t), \quad B \in \mathcal{B}(\mathbb{R}).$$

Actually, this means that the distribution $\mu_X \ll \lambda_1$, where the Lebesgue measure λ_1 is considered on $\mathcal{B}(\mathbb{R})$, and moreover the Radon–Nikodym derivative $\frac{d\mu_X}{d\lambda_1} = f(t) \pmod{\lambda_1}$.

DEFINITION 1.5.– A r.v. γ is called normal (or normally distributed) if it has a pdf of the form

$$\rho(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-m)^2}{2\sigma^2}}, \quad x \in \mathbb{R},$$

with parameters $m \in \mathbb{R}$ and $\sigma > 0$. This is denoted as follows: $\gamma \sim N(m, \sigma^2)$.

If $\gamma \sim N(m, \sigma^2)$, then it holds

$$E\gamma = m, \quad D\gamma = \sigma^2.$$

Hereafter, E stands for expectation and D stands for the variance of a r.v. Remember that

$$E\gamma = \int_{\Omega} \gamma(\omega) dP(\omega), \quad D\gamma = E(\gamma - m)^2 = \int_{\Omega} (\gamma(\omega) - m)^2 dP(\omega).$$

By the change of variables formula (theorem 1.2), it holds

$$E\gamma = \int_{\mathbb{R}} x d\mu_{\gamma}(x), \quad D\gamma = \int_{\mathbb{R}} (x - m)^2 d\mu_{\gamma}(x),$$

and since γ has pdf equal ρ , we have

$$m = E\gamma = \int_{\mathbb{R}} x\rho(x)dx, \quad \sigma^2 = D\gamma = \int_{\mathbb{R}} (x - m)^2 \rho(x)dx.$$

The latter integrals are Lebesgue integrals.

DEFINITION 1.6.—A r.v. γ has degenerate normal distribution $N(m, 0)$ if $\gamma(\omega) = m$, almost surely (a.s.). We denote it as $\gamma \sim N(m, 0)$.

In this case, $E\gamma = m$, $D\gamma = 0$, and the distribution μ_γ is a point measure concentrated at m :

$$\mu_\gamma(B) = I_B(m), \quad B \in \mathcal{B}(\mathbb{R}).$$

Such a measure is called Dirac measure at point m and is denoted as δ_m .

DEFINITION 1.7.—A r.v. γ is called Gaussian if it is either normally distributed (with positive variance) or has degenerate normal distribution (with zero variance).

Thus, a Gaussian r.v. γ satisfies $\gamma \sim N(m, \sigma^2)$ with some $m \in \mathbb{R}$ and $\sigma \geq 0$. If $\sigma > 0$, then $\mu_\gamma \ll \lambda_1$, and if $\sigma = 0$, then $\mu_\gamma = \delta_m \perp \lambda_1$ (i.e. μ_γ is singular to λ_1).

DEFINITION 1.8.—A r.v. $\gamma \sim N(0, 1)$ is called standard normal.

Remember that characteristic function φ_ξ of a r.v. ξ is as follows:

$$\varphi_\xi(t) = E e^{it\xi}, \quad t \in \mathbb{R}.$$

A normal r.v. $\gamma \sim N(m, \sigma^2)$ has characteristic function

$$\varphi_\gamma(t) = \exp\left\{imt - \frac{\sigma^2 t^2}{2}\right\}. \quad [1.17]$$

If γ has degenerate normal distribution $N(m, 0)$, then

$$\varphi_\gamma(t) = \exp\{imt\}.$$

Thus, relation [1.17] holds true for any Gaussian r.v. $\gamma \sim N(m, \sigma^2)$, with $\sigma \geq 0$.

1.4.2. Random vectors

Remember that a random vector X distributed in \mathbb{R}^n is \mathcal{F} -measurable mapping $X : \Omega \rightarrow \mathbb{R}^n$, where (Ω, \mathcal{F}, P) is the underlying probability space.

A mapping $X(\omega) = (X_1(\omega), \dots, X_n(\omega))^\top$, $\omega \in \Omega$ is a random vector if, and only if, all $X_k(\omega)$, $k = 1, \dots, n$ are random variables.

For a random vector X , its expectation is defined coordinate-wise:

$$E X = (E X_1, \dots, E X_n)^\top =: m, \quad m = (m_1, \dots, m_n)^\top \in \mathbb{R}^n.$$

In case $E X_k^2 < \infty$ for all $k = 1, \dots, n$, its variance–covariance matrix $\text{Cov}(X) = (s_{ij})_{i,j=1}^n$ is defined as follows:

$$s_{ij} = \text{Cov}(X_i, X_j) = E(X_i - m_i)(X_j - m_j), \quad i, j = 1, \dots, n. \quad [1.18]$$

Hereafter, expectation E is considered as an operator that acts on the total product under its sign, and we omit brackets for brevity.

The variance–covariance matrix can be expressed as

$$\text{Cov}(X) = E(X - m)(X - m)^\top. \quad [1.19]$$

Here, the expectation of a random matrix is a matrix composed of expectations of entries, according to [1.18]. The variance–covariance matrix is a positive semidefinite matrix, i.e. it is symmetric and the corresponding quadratic form is non-negative. The variance–covariance matrix of a random vector X exists if and only if $E \|X\|^2 < \infty$.

LEMMA 1.2.– (Variance–covariance matrix under linear transform) Let X be a random vector in \mathbb{R}^n , with variance–covariance matrix S , and $A \in \mathbb{R}^{p \times n}$. Then AX is a random vector in \mathbb{R}^p , with

$$\text{Cov}(AX) = ASA^\top. \quad [1.20]$$

PROOF.– The mapping $AX : \Omega \rightarrow \mathbb{R}^p$ is \mathcal{F} -measurable, as a Borel function of random vector. Hence AX is a random vector in \mathbb{R}^p , and because $E \|AX\|^2 \leq \|A\|^2 \cdot E \|X\|^2 < \infty$, it has a variance–covariance matrix.

Hereafter, $\|A\|$ is Euclidean norm of matrix A ,

$$\|A\| = \sup_{x \neq 0} \frac{\|Ax\|}{\|x\|}.$$

Now, $E(AX) = A(E X) = Am$, where $m = E X$, and

$$\begin{aligned} \text{Cov}(AX) &= E(AX - Am)(AX - Am)^\top = \\ &= E[A(X - m)(X - m)^\top A^\top] = A \cdot E(X - m)(X - m)^\top \cdot A^\top = ASA^\top. \end{aligned}$$

Here, we used the linearity of the operator of taking expectation. □

COROLLARY 1.3.– (Moments of linear functional) Let $a \in \mathbb{R}^n$ and X be a random vector in \mathbb{R}^n , with mean value m and variance–covariance matrix S . Then

$$E(a^\top X) = a^\top m, \quad D(a^\top X) = a^\top S a.$$

PROOF.– The statement follows from lemma 1.2 and its proof if to put $A = a^\top \in \mathbb{R}^{1 \times n}$. The random vector $a^\top X$ is just a r.v., and its variance–covariance matrix is just the variance. □

1.4.3. Distributions of random vectors

Let X be a random vector distributed in \mathbb{R}^n . Its distribution is introduced similarly to definition 1.4.

DEFINITION 1.9.– *The distribution of X is a probability measure μ_X defined as follows:*

$$\mu_X(B) = P\{\omega : X(\omega) \in B\}, \quad B \in \mathcal{B}(\mathbb{R}^n).$$

It is always possible to construct a random vector with a given distribution.

LEMMA 1.3.– Given a probability measure μ on $\mathcal{B}(\mathbb{R}^n)$, there exists a random vector X , with distribution $\mu_X = \mu$.

PROOF.– Take the measure space $(\mathbb{R}^n, \mathcal{B}(\mathbb{R}^n), \mu)$ as a probability space (Ω, \mathcal{F}, P) and define $X : \Omega \rightarrow \mathbb{R}^n$ as $X(\omega) = \omega$. Then

$$\mu_X(B) = P\{\omega \in B\} = P(B) = \mu(B), \quad B \in \mathcal{B}(\mathbb{R}^n). \quad \square$$

Remember that random variables X_1, \dots, X_n , which are defined on the same probability space, are independent if

$$P\{X_1 \in B_1, X_2 \in B_2, \dots, X_n \in B_n\} = \prod_1^n P\{X_k \in B_k\},$$

for all $B_1, \dots, B_n \in \mathcal{B}(\mathbb{R})$.

The latter relation can be written in terms of the distribution μ_X of random vector $X = (X_k)_1^n$ and marginal distributions μ_{X_k} of its components:

$$\mu_X\left(\prod_1^n B_k\right) = \prod_1^n \mu_{X_k}(B_k).$$

Here $\prod_1^n B_k$ denotes Cartesian product of the sets B_k .

It is clear that components of random vector $X = (X_k)_1^n$ are independent if, and only if, μ_X is a product of n probability measures, and in this case

$$\mu_X = \prod_1^n \mu_{X_k}.$$

Remember that characteristic function φ_X of a random vector X is defined as follows:

$$\varphi_X(t) = E e^{i(X,t)}, \quad t \in \mathbb{R}^n.$$

One can rewrite $\varphi_X(t)$ using the change of variables formula (see theorem 1.2):

$$\varphi_X(t) = \int_{\Omega} e^{i(X(\omega), t)} dP(\omega) = \int_{\mathbb{R}^n} e^{i(z, t)} d(PX^{-1})(z),$$

$$\varphi_X(t) = \int_{\mathbb{R}^n} e^{i(z, t)} d\mu_X(z).$$

This prompts the following definition.

DEFINITION 1.10.– Given a probability measure μ on $\mathcal{B}(\mathbb{R}^n)$, its characteristic function φ_{μ} is as follows:

$$\varphi_{\mu}(t) = \int_{\mathbb{R}^n} e^{i(z, t)} d\mu(z), t \in \mathbb{R}^n.$$

Thus, φ_X and φ_{μ_X} coincide.

From standard course of probability theory, it is known that the cumulative distribution function (and therefore, the distribution) of a random vector is uniquely defined by its characteristic function.

LEMMA 1.4.– (Criterion for independence) Consider a random vector $X = (X_k)_1^n$. Its components are independent if, and only if, φ_X can be decomposed as follows:

$$\varphi_X(t) = \varphi_1(t_1)\varphi_2(t_2)\dots\varphi_n(t_n), t \in \mathbb{R}^n,$$

where $\varphi_k : \mathbb{R} \rightarrow \mathbb{C}$ are some functions with $\varphi_k(0) = 1, k = 1, \dots, n$, and in this case

$$\varphi_X(t) = \prod_1^n \varphi_{X_k}(t_k), t \in \mathbb{R}^n.$$

PROOF.–

a) Let X_k be independent. Then random variables $e^{it_k X_k}, k = 1, \dots, n$ are independent as well, and

$$\varphi_X(t) = \mathbb{E} \prod_1^n e^{it_k X_k} = \prod_1^n \mathbb{E} e^{it_k X_k} = \prod_1^n \varphi_{X_k}(t_k); \quad \varphi_{X_k}(0) = 1.$$

b) Assume that $\varphi_X(t) = \varphi_1(t_1)\varphi_2(t_2)\dots\varphi_n(t_n)$, with $\varphi_k(0) = 1, k = 1, \dots, n$. Let $\{e_k, k = 1, \dots, n\}$ be the standard orthobasis in \mathbb{R}^n . Then

$$\varphi_{X_k}(t_k) = \mathbb{E} e^{it_k X_k} = \mathbb{E} e^{i(X, t_k e_k)} = \varphi_X(t_k e_k) = \varphi_k(t_k).$$

Let $Y = (Y_k)_1^n$ be a random vector with independent components and the same marginal distributions:

$$\mu_{Y_k} = \mu_{X_k}, \quad k = 1, \dots, n.$$

(Such Y can be constructed if to apply lemma 1.3 to the measure $\mu = \prod_1^n \mu_{X_k}$.) Then, by part (a) of the proof,

$$\varphi_Y(t) = \prod_1^n \varphi_{Y_k}(t_k) = \prod_1^n \varphi_{X_k}(t_k) = \varphi_X(t), \quad t \in \mathbb{R}^n.$$

Therefore,

$$\mu_X = \mu_Y = \prod_1^n \mu_{Y_k} = \prod_1^n \mu_{X_k},$$

and the components of X are independent. \square

In an obvious way, lemma 1.3 can be reformulated as a criterion for a probability measure on $\mathcal{B}(\mathbb{R}^n)$ to be a product of n probability measures on $\mathcal{B}(\mathbb{R})$.

DEFINITION 1.11.— Given a probability measure μ on $\mathcal{B}(\mathbb{R}^n)$, its mean value m_μ and variance–covariance matrix $\text{Cov}(\mu) = (s_{ij})_{i,j=1}^n$ are defined as follows:

$$m_\mu = \int_{\mathbb{R}^n} x d\mu(x) := \left(\int_{\mathbb{R}^n} x_k d\mu(x) \right)_{k=1}^n = (m_k)_1^n,$$

$$s_{ij} = \int_{\mathbb{R}^n} (x_i - m_i)(x_j - m_j) d\mu(x), \quad i, j = 1, \dots, n.$$

Definition 1.11 is consistent with the corresponding definition of the mean and variance–covariance matrix of a random vector. Indeed, for a random vector X , it holds

$$\mathbb{E} X = m_{\mu_X}, \quad \text{Cov}(X) = \text{Cov}(\mu_X),$$

i.e. expectation and variance–covariance matrix of a random vector are just the mean and variance–covariance matrix of its distribution.

Now, we interpret the bilinear form generated by $S = \text{Cov}(X)$. Let $m = \mathbb{E} X$ and $u, v \in \mathbb{R}^n$; then

$$(Su, v) = v^\top (\mathbb{E}(X - m)(X - m)^\top) u = \mathbb{E} [v^\top (X - m)(X - m)^\top u],$$

$$(Su, v) = \mathbb{E}(X - m, u)(X - m, v).$$

For a probability measure μ on $\mathcal{B}(\mathbb{R}^n)$, with $m = m_\mu$ and $S = \text{Cov}(\mu)$, we have, respectively:

$$(Su, v) = \int_{\mathbb{R}^n} (z - m, u)(z - m, v) d\mu(z), \quad u, v \in \mathbb{R}^n.$$

For the mean value, we have

$$(m, u) = \int_{\mathbb{R}^n} (z, u) d\mu(z), \quad u \in \mathbb{R}^n.$$

Those expressions are the first and the central second moments of measure μ .

Problems 1.4

20) A measure μ on $\mathcal{B}(\mathbb{R}^n)$ is called symmetric around the origin if $\mu(B) = \mu(-B)$, for all $B \in \mathcal{B}(\mathbb{R}^n)$. Let μ be a probability measure on $\mathcal{B}(\mathbb{R}^n)$. Prove that μ is symmetric around the origin if, and only if, its characteristic function φ_μ takes real values only.

21) Let μ be a probability measure on $\mathcal{B}(\mathbb{R}^n)$. Prove that μ is invariant under all orthogonal transformations if, and only if, there exists a function $f : [0, +\infty) \rightarrow \mathbb{C}$ such that $\varphi_\mu(t) = f(\|t\|)$, $t \in \mathbb{R}^n$.

22) Let μ and ν be probability measures on $\mathcal{B}(\mathbb{R}^n)$ such that $\mu(\bar{B}(x, r)) = \nu(\bar{B}(x, r))$, for all closed balls $\bar{B}(x, r)$ in \mathbb{R}^n . Prove that $\mu = \nu$.

1.5. Gaussian vectors and Gaussian measures

1.5.1. Characteristic functions of Gaussian vectors

DEFINITION 1.12.– A random vector ξ in \mathbb{R}^n is called Gaussian if for each $a \in \mathbb{R}^n$, inner product (ξ, a) is a Gaussian r.v.

Consider a Gaussian random vector $\xi = (\xi_k)_1^n$ in \mathbb{R}^n . Its components $\xi_k = (\xi, e_k)$ are Gaussian random variables, $\xi_k \sim N(m_k, \sigma_k^2)$ with $\sigma_k \geq 0$; $k = 1, \dots, n$. Such random variables, the components of a Gaussian random vector, are called jointly Gaussian. It holds

$$\mathbb{E} \|\xi\|^2 = \sum_1^n \mathbb{E} \xi_k^2 = \sum_1^n (m_k^2 + \sigma_k^2) < \infty,$$

and therefore, $\text{Cov}(\xi)$ is well defined. We have

$$\begin{aligned} \mathbf{E} \xi &= (\mathbf{E} \xi_k)_1^n = (m_k)_1^n =: m, \quad m \in \mathbb{R}^n; \\ S &:= \text{Cov}(\xi) = (s_{ij})_{i,j=1}^n, \quad s_{ij} = \text{Cov}(\xi_i, \xi_j), \\ s_{ii} &= \mathbf{D} \xi_i = \sigma_i^2, \quad 1 \leq i, j \leq n. \end{aligned}$$

Then we write

$$\xi \sim N(m, S)$$

and say that ξ is a Gaussian random vector with mean m and variance–covariance matrix S . Here, S is a positive semidefinite $n \times n$ real matrix as a variance–covariance matrix of a random vector in \mathbb{R}^n .

LEMMA 1.5.– If $\xi \sim N(m, S)$ then for each $a \in \mathbb{R}^n$,

$$(\xi, a) \sim N(m_a, \sigma_a^2), \quad m_a = (m, a), \quad \sigma_a^2 = (Sa, a).$$

PROOF.– R.v. (ξ, a) is Gaussian according to definition 1.12. Its mean and variance are evaluated in corollary 1.3. \square

LEMMA 1.6.– If $\xi \sim N(m, S)$ in \mathbb{R}^n , then

$$\varphi_\xi(t) = \exp \left\{ i(t, m) - \frac{(St, t)}{2} \right\}, \quad t \in \mathbb{R}^n.$$

PROOF.– It holds

$$\varphi_\xi(t) = \mathbf{E} e^{i(\xi, t)} = \varphi_{(\xi, t)}(1).$$

Now, use lemmas 1.5 and [1.17]

$$\varphi_\xi(t) = \exp \left\{ iz(m, t) - \frac{z^2(St, t)}{2} \right\} \Big|_{z=1},$$

and the statement follows. \square

REMARK 1.3.– If a random vector ξ has characteristic function given in lemma 1.6, with certain $n \times n$ real and symmetric matrix S , then S is positive semidefinite and $\xi \sim N(m, S)$.

PROOF.– For $a \in \mathbb{R}^n$, it holds

$$\begin{aligned} \varphi_{(\xi, a)}(u) &= \mathbf{E} e^{i(\xi, ua)} = \varphi_\xi(ua), \\ \varphi_{(\xi, a)}(u) &= \exp \left\{ iu(a, m) - \frac{(Sa, a)u^2}{2} \right\}, \quad u \in \mathbb{R}. \end{aligned}$$

Since the absolute value of characteristic function does not exceed 1, matrix S is positive semidefinite, and moreover $(\xi, a) \sim N(m_a, \sigma_a^2)$, with $m_a = (a, m)$, $\sigma_a^2 = (Sa, a)$. Hence, ξ is a Gaussian vector, with some parameters $m_1 \in \mathbb{R}^n$ and $S_1 \in \mathbb{R}^{n \times n}$, S_1 is positive semidefinite. Then lemma 1.5 implies $(\xi, a) \sim N(\tilde{m}_a, \tilde{\sigma}_a^2)$, with $\tilde{m}_a = (a, m_1)$, $\tilde{\sigma}_a^2 = (S_1 a, a)$. We get $(a, m) = (a, m_1)$ and $(Sa, a) = (S_1 a, a)$ for all $a \in \mathbb{R}^n$. Thus, $m_1 = m$ and $S_1 = S$. \square

DEFINITION 1.13.— *Random vector $\gamma \sim N(0, I_n)$ is called standard, or canonical, Gaussian vector distributed in \mathbb{R}^n .*

The components $\gamma_1, \dots, \gamma_n$ of standard Gaussian vector are uncorrelated jointly Gaussian standard random variables. In particular, $E \gamma_k = 0$, $D \gamma_k = 1$, and for all $i \neq j$, $\text{Cov}(\gamma_i, \gamma_j) = E \gamma_i \gamma_j = 0$.

THEOREM 1.6.— (About components of standard Gaussian vector)

1) Let $\gamma = (\gamma_k)_1^n$ be standard Gaussian vector. Then

$$\varphi_\gamma(t) = e^{-\frac{\|t\|^2}{2}}, \quad t \in \mathbb{R}^n$$

and $\gamma_1, \dots, \gamma_n$ are independent and identically distributed (i.i.d.) $N(0, 1)$ random variables.

2) If $\gamma_1, \dots, \gamma_n$ are i.i.d. $N(0, 1)$ random variables, then $\gamma = (\gamma_k)_1^n$ is standard Gaussian vector.

PROOF.—

1) The formula for φ_γ follows from lemma 1.5 with $m = 0$ and $S = I_n$. Therefore,

$$\varphi_\gamma(t) = \prod_{k=1}^n e^{-\frac{t_k^2}{2}}.$$

Now, by lemma 1.3 the components of γ are independent.

2) Let $\gamma_1, \dots, \gamma_n$ be i.i.d. $N(0, 1)$ random variables. Then $\gamma = (\gamma_k)_1^n$ is a random vector, and for each $a \in \mathbb{R}^n$, $(\gamma, a) = \sum_{k=1}^n a_k \gamma_k$ is a Gaussian r.v. as a sum of independent Gaussian random variables. Therefore, γ is Gaussian. We have

$$E \gamma = (E \gamma_k)_1^n = 0,$$

$$\text{Cov } \gamma = (s_{ij})_{i,j=1}^n, \quad s_{ij} = E \gamma_i \gamma_j = \delta_{ij}, \quad 1 \leq i, j \leq n.$$

Hereafter, δ_{ij} is Kronecker delta, $\delta_{ij} = 1$ if $i = j$ and $\delta_{ij} = 0$, otherwise. Hence $\gamma \sim N(0, I_n)$. \square

REMARK 1.4.— (About uncorrelated Gaussian variables) Theorem 1.6 can be extended as follows: jointly Gaussian random variables ξ_1, \dots, ξ_n are independent if, and only if, they are uncorrelated. Based on theorem 1.6, this can be proven by consideration of normalized random variables $\eta_i = \frac{\xi_i - E \xi_i}{\sqrt{D \xi_i}}$ (before we cancel that ξ_i are constant a.s.).

1.5.2. Expansion of Gaussian vector

Lemma 1.5 can be generalized for any linear transformation of a Gaussian vector.

LEMMA 1.7.– (Linear transform of Gaussian vector) Let $\xi \sim N(m, S)$ be a Gaussian vector in \mathbb{R}^n and $A \in \mathbb{R}^{p \times n}$. Then $A\xi$ is a Gaussian vector in \mathbb{R}^p and

$$A\xi \sim N(Am, ASA^\top).$$

PROOF.– For $a \in \mathbb{R}^p$, it holds $(A\xi, a) = (\xi, A^\top a)$. It is a Gaussian r.v. due to definition 1.12. Now, the statement follows from lemma 1.2 and its proof. \square

Now, we want to show that for any $m \in \mathbb{R}^n$ and positive semidefinite matrix $S \in \mathbb{R}^{n \times n}$, a random vector $\xi \sim N(m, S)$ can be obtained as an affine transformation of standard Gaussian vector.

Let $A \in \mathbb{R}^{n \times n}$ be a positive semidefinite matrix. Then it has orthonormal eigenbasis e_1, \dots, e_n and corresponding non-negative eigenvalues $\lambda_1, \dots, \lambda_n$. The matrix can be expanded as follows:

$$A = \sum_1^n \lambda_k e_k e_k^\top.$$

DEFINITION 1.14.– Let A be a positive semidefinite matrix as described above. $A^{\frac{1}{2}} = \sqrt{A}$ is a matrix satisfying

$$\sqrt{A} = \sum_1^n \sqrt{\lambda_k} e_k e_k^\top.$$

It is a unique positive semidefinite matrix with its square equal to A . The matrix \sqrt{A} has the same eigenbasis e_1, \dots, e_n and corresponding eigenvalues $\sqrt{\lambda_1}, \dots, \sqrt{\lambda_n}$.

LEMMA 1.8.– (Representation of Gaussian vector via standard Gaussian vector) Let $m \in \mathbb{R}^n$, $S \in \mathbb{R}^{n \times n}$ be a positive semidefinite matrix and $\gamma \sim N(0, I_n)$. Then

$$\xi := m + \sqrt{S}\gamma \sim N(m, S).$$

PROOF.– According to lemma 1.7

$$\sqrt{S}\gamma \sim N(0, \sqrt{S}I_n(\sqrt{S})^\top), \quad \sqrt{S}\gamma \sim N(0, S).$$

This implies the statement. \square

Since $\gamma \sim N(0, I_n)$ can be constructed based on i.i.d. $N(0, 1)$ random variables, lemma 1.8 shows that Gaussian $N(m, S)$ random vectors do exist, for any $m \in \mathbb{R}^n$ and any positive semidefinite matrix $S \in \mathbb{R}^{n \times n}$.

Now, we expand a Gaussian vector $\xi \sim N(m, S)$ using i.i.d. $N(0, 1)$ random variables $\gamma_1, \dots, \gamma_r$, with $r = rk(S)$.

THEOREM 1.7.– (Decomposition of Gaussian vector) Let $m \in \mathbb{R}^n$ and $S \in \mathbb{R}^{n \times n}$ be a positive semidefinite matrix, with $rk(S) = r \geq 1$. Let $\lambda_1, \dots, \lambda_r$ be positive eigenvalues of S and e_1, \dots, e_r be the corresponding orthonormal system of eigenvectors.

1) For $\xi \sim N(m, S)$, there exist i.i.d. $N(0, 1)$ random variables $\gamma_1, \dots, \gamma_r$ on the underlying probability space such that

$$\xi = m + \sum_1^r \sqrt{\lambda_k} \gamma_k e_k, \quad \text{a.s.}$$

2) If $\gamma_1, \dots, \gamma_r$ are i.i.d. $N(0, 1)$ random variables, then

$$\eta := m + \sum_1^r \sqrt{\lambda_k} \gamma_k e_k \sim N(m, S).$$

PROOF.– We complete the orthonormal system to an orthobasis e_1, \dots, e_n . Denote $\lambda_{r+1} = \lambda_{r+2} = \dots = \lambda_n = 0$; they are the rest eigenvalues of S .

1) Introduce $X = \xi - m$, $X \sim N(0, S)$. Now, $X = \sum_1^n (X, e_k) e_k$, $E(X, e_k) = 0$, and for all $k \geq r + 1$, $D(X, e_k) = \lambda_k = 0$. Hence $(X, e_k) = 0$ a.s., for all $k \geq r + 1$. Thus,

$$X = \sum_1^r \sqrt{\lambda_k} \frac{(X, e_k)}{\sqrt{\lambda_k}} e_k, \quad \text{a.s.}$$

Random variables $\gamma_k := \frac{(X, e_k)}{\sqrt{\lambda_k}}$, $k = 1, \dots, r$ are jointly Gaussian (because vector $\gamma = (\gamma_k)_1^r$ is a linear transformation of Gaussian vector X) and $E \gamma_k = 0$,

$$\text{Cov}(\gamma_k, \gamma_j) = \frac{1}{\sqrt{\lambda_k \lambda_j}} (S e_k, e_j) = \frac{\lambda_k}{\sqrt{\lambda_k \lambda_j}} \delta_{kj} = \delta_{kj}.$$

According to theorem 1.6(1), $\gamma_1, \dots, \gamma_r$ are i.i.d. $N(0, 1)$ random variables. Now,

$$X = \sum_1^r \sqrt{\lambda_k} \gamma_k e_k, \quad \text{a.s.}$$

and the statement follows.

2) For $a \in \mathbb{R}^n$, $(\eta, a) = (m, a) + \sum_{k=1}^r \sqrt{\lambda_k} \gamma_k (e_k, a)$ is a Gaussian r.v. as a sum of independent Gaussian random variables. Therefore, η is Gaussian. Next,

$$\begin{aligned} E\eta &= m, \quad \text{Cov}(\eta) = E \left(\sum_{k=1}^r \sqrt{\lambda_k} \gamma_k e_k \right) \left(\sum_{j=1}^r \sqrt{\lambda_j} \gamma_j e_j^\top \right) = \\ &= \sum_{k,j=1}^r \sqrt{\lambda_k \lambda_j} (E \gamma_k \gamma_j) e_k e_j^\top = \sum_{k=1}^r \lambda_k e_k e_k^\top = S. \end{aligned}$$

Thus, $\eta \sim N(m, S)$. □

1.5.3. Support of Gaussian vector

DEFINITION 1.15.— For a random vector X in \mathbb{R}^n , denote by G a union of all balls $B(x, r)$, with $P\{X \in B(x, r)\} = 0$. The set $\mathbb{R}^n \setminus G$ is called support of X and denoted as $\text{supp } X$.

It is clear that $P\{X \in G\} = 0$. Moreover, G is the largest open set with this property. $\text{supp } X$ is the smallest closed set such that X belongs to this set with probability 1. Since $G \neq \mathbb{R}^n$, $\text{supp } X \neq \emptyset$.

EXAMPLE 1.2.— Let that a r.v. ξ has Poisson distribution with parameter $\lambda > 0$. Then $\text{supp } \xi = \{0, 1, 2, \dots, n, \dots\}$.

THEOREM 1.8.— (About support of Gaussian vector) For random vector $\xi \sim N(0, S)$, $\text{supp } \xi = R(S)$, where $R(S)$ is the range of S .

PROOF.—

a) If $S = 0$, then $\xi = 0$ a.s., and $\text{supp } \xi = \{0\} = R(S)$.

Now, let $r_k(S) = r \geq 1$, and let $\lambda_1, \dots, \lambda_r$ be positive eigenvalues of S and e_1, \dots, e_r be the corresponding orthonormal system of eigenvectors. According to theorem 1.7(1), there exist i.i.d. $N(0, 1)$ random variables $\gamma_1, \dots, \gamma_r$ on the underlying probability space such that

$$\xi = \sum_{k=1}^r \sqrt{\lambda_k} \gamma_k e_k, \quad \text{a.s.},$$

and with probability one $\xi \in \text{span}(e_1, \dots, e_r) = R(S)$.

b) Take arbitrary $x \in R(S)$ and $\varepsilon > 0$. Show that

$$P\{\xi \in B(x, \varepsilon)\} > 0. \tag{1.21}$$

Indeed, $x = \sum_1^r a_k e_k$ and there exists $\delta > 0$ such that

$$M := \left\{ y = \sum_1^r b_k e_k : |b_k - a_k| < \delta, k = 1, \dots, r \right\} \subset B(x, \varepsilon).$$

Then

$$\begin{aligned} \mathbb{P}\{\xi \in M\} &= \mathbb{P}\left\{ \gamma_k \in \left(\frac{a_k - \delta}{\sqrt{\lambda_k}}, \frac{a_k + \delta}{\sqrt{\lambda_k}} \right), k = 1, \dots, r \right\} = \\ &= \prod_1^r \mathbb{P}\left\{ \gamma_k \in \left(\frac{a_k - \delta}{\sqrt{\lambda_k}}, \frac{a_k + \delta}{\sqrt{\lambda_k}} \right) \right\} > 0, \\ \mathbb{P}\{\xi \in B(x, \varepsilon)\} &\geq \mathbb{P}\{\xi \in M\} > 0. \end{aligned}$$

Thus, [1.21] holds true. This fact and the relation shown in part (a) imply that $\text{supp } \xi = R(S)$. \square

1.5.4. Gaussian measures in Euclidean space

Now, we reformulate the results of sections 1.5.1–1.5.3 for distributions of Gaussian vectors.

DEFINITION 1.16.—A probability measure μ on $\mathcal{B}(\mathbb{R}^n)$ is called *Gaussian measure in \mathbb{R}^n* , if there exists a Gaussian random vector ξ in \mathbb{R}^n such that its distribution $\mu_\xi = \mu$.

Let $\xi \sim N(m, S)$ and $\mu = \mu_\xi$. In view of section 1.4.3, we have the following:

$$m = \mathbb{E} \xi = m_\mu, \quad S = \text{Cov}(\xi) = \text{Cov}(\mu),$$

i.e. m is mean value of μ and S is variance–covariance matrix of μ .

Next, according to lemma 1.6 the characteristic function of μ equals

$$\varphi_\mu(t) = \varphi_\xi(t) = \exp \left\{ i(t, m) - \frac{(St, t)}{2} \right\}, \quad t \in \mathbb{R}^n.$$

Since parameters μ and S define uniquely the characteristic function of a Gaussian measure, they define uniquely the Gaussian measure μ itself. There is one-to-one correspondence between the set of all Gaussian measures in \mathbb{R}^n and the set of couples $(m; S)$, where $m \in \mathbb{R}^n$ and S is a positive semidefinite $n \times n$ matrix.

Let $\lambda_1, \dots, \lambda_n$ be eigenvalues of S (they are non-negative) and e_1, \dots, e_n be the corresponding eigenbasis of S . Then

$$\varphi_\mu(t) = \exp \left\{ i \sum_1^n (t, e_k)(m, e_k) - \frac{1}{2} \sum_1^n \lambda_k (t, e_k)^2 \right\},$$

$$\varphi_\mu(t) = \prod_1^n \exp \left\{ i(t, e_k)(m, e_k) - \frac{1}{2} \lambda_k(t, e_k)^2 \right\}.$$

We treat (t, e_k) as coordinates t_k of vector $t \in \mathbb{R}^n$, and the same for $(m, e_k) = m_k$. Thus,

$$\varphi_\mu(t) = \prod_1^n \exp \left\{ it_k m_k - \frac{1}{2} \lambda_k t_k^2 \right\}.$$

In view of lemma 1.3, we get a decomposition of μ :

$$\mu = \prod_1^n \mu_k,$$

where μ_k is a Gaussian measure on real line, with mean m_k and variance λ_k , $k = 1, \dots, n$. Thus, each Gaussian measure in Euclidean space is just a product of Gaussian measures on real line.

In case, $rk(S) = r$, $1 \leq r \leq n - 1$ we may and do assume that $\lambda_{r+1} = \lambda_{r+2} = \dots = \lambda_n = 0$. Then $\mu_k = \delta_{m_k}$ (Dirac measure at point m_k), $k \geq r + 1$, and we get the expansion

$$\mu = \prod_1^r \mu_k \times \prod_{r+1}^n \delta_{m_k} = \prod_1^r \mu_k \times \delta_z, \quad z = (m_{r+1}, \dots, m_n)^\top.$$

Here, δ_z is Dirac measure on $\mathcal{B}(\mathbb{R}^{n-r})$ at point z .

DEFINITION 1.17.— *Standard Gaussian measure g in \mathbb{R}^n is the distribution of standard Gaussian vector in \mathbb{R}^n .*

The measure g has zero mean and the variance–covariance matrix is equal to I_n . Its characteristic function is

$$\varphi_g(t) = e^{-\frac{\|t\|^2}{2}}, t \in \mathbb{R}^n.$$

It holds

$$g = \prod_1^n g_k,$$

where $g_1 = \dots = g_n$ is standard Gaussian measure on real line. This means that

$$g_i(B) = \int_B \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx, \quad B \in \mathcal{B}(\mathbb{R}), \quad i = 1, \dots, n.$$

DEFINITION 1.18.— *(See definition 1.15) For a probability measure μ on $\mathcal{B}(\mathbb{R}^n)$, denote by G a union of all balls $B(x, r)$, with $\mu(B(x, r)) = 0$. The set $\mathbb{R}^n \setminus G$ is called support of μ and is denoted as $\text{supp } \mu$.*

Since \mathbb{R}^n is separable, G is a countable union of balls $B(x_i, r_i)$, with $\mu(B(x_i, r_i)) = 0$. Hence, $\mu(G) = 0$. Moreover, G is the largest open set with this property. Note that $\text{supp } \mu$ is the smallest closed set such that the value of μ at this set equals 1. Since $G \neq \mathbb{R}^n$, $\text{supp } \mu \neq \emptyset$.

Theorem 1.8 implies the following: if μ is a Gaussian measure with mean 0 and variance–covariance matrix S , then $\text{supp } \mu = R(S)$. In particular for standard Gaussian measure g in \mathbb{R}^n , $\text{supp } g = \mathbb{R}^n$.

Now, we study the invariance of Gaussian measures under linear transformations. Let X be a random vector in \mathbb{R}^n with distribution μ_X and $T : \mathbb{R}^n \rightarrow \mathbb{R}^n$ be a Borel function. Random vector TX has distribution $\mu_{TX} = \mu_X T^{-1}$. Therefore, μ_X is invariant under T (see definition 1.2) if, and only if, $X \stackrel{d}{=} TX$. Hereafter, $X \stackrel{d}{=} Y$ means that random vectors X and Y are identically distributed, i.e. $\mu_X = \mu_Y$.

Remember definition 1.14 of $A^{1/2} = \sqrt{A}$ for a positive semidefinite matrix A .

THEOREM 1.9.– (Invariance of Gaussian measure) Let $U \in \mathbb{R}^{n \times n}$ and μ be a Gaussian measure in \mathbb{R}^n , with zero mean and non-singular variance–covariance matrix S . The measure μ is U -invariant if, and only if, the matrix $S^{-1/2}US^{1/2}$ is orthogonal. In particular, standard Gaussian measure g in \mathbb{R}^n is U -invariant if, and only if, U is an orthogonal matrix.

PROOF.– Let X be a Gaussian random vector, with $\mu_X = \mu$. Then $X \sim N(0, S)$ and by lemma 1.7, $UX \sim N(m, USU^T)$. Now, $UX \stackrel{d}{=} X$ if, and only if,

$$USU^T = S \iff \left(S^{-1/2}US^{1/2}\right) \left(S^{-1/2}US^{1/2}\right)^T = I_n.$$

But this is equivalent to the orthogonality of matrix $S^{-1/2}US^{1/2}$.

In case $\mu = g$, it holds $S = I_n$. Thus, g is U -invariant if, and only if, U is an orthogonal matrix. □

The main statement of theorem 1.9 can be interpreted as follows. For a positive definite $S \in \mathbb{R}^{n \times n}$, we introduce a new inner product in \mathbb{R}^n ,

$$(x, y)_S = (S^{-1}x, y), \quad x, y \in \mathbb{R}^n.$$

The corresponding norm is

$$\|x\|_S = \sqrt{(x, x)_S} = \sqrt{\|S^{-\frac{1}{2}}x\|^2} = \|S^{-\frac{1}{2}}x\|, \quad x \in \mathbb{R}^n.$$

Now, a Gaussian measure μ , with zero mean and non-singular variance–covariance matrix S , is U -invariant if, and only if, U is unitary operator w.r.t. the inner product $(x, y)_S$.

Indeed, $S^{-1/2}US^{1/2}$ is an orthogonal matrix if, and only if,

$$\|S^{-1/2}US^{1/2}x\| = \|x\|, \quad x \in \mathbb{R}^n.$$

Now, make a change of variable $y = S^{1/2}x$, $y \in \mathbb{R}^n$. Then we get an equivalent condition

$$\|S^{-1/2}Uy\| = \|S^{-1/2}y\| \Leftrightarrow \|Uy\|_S = \|y\|_S, \quad y \in \mathbb{R}^n.$$

The latter equality means that the linear transformation U is orthogonal w.r.t. the inner product $(x, y)_S$. One can say that a Gaussian measure μ changes the geometry of Euclidean space. Standard Gaussian measure g stays in correspondence with standard geometry of Euclidean space. Compared to Lebesgue measure λ_n (see theorem 1.3 and corollary 1.1), the measure g has fewer invariant transformations. Theorem 1.6 shows that λ_n cannot be extended to an infinite-dimensional Hilbert space H . But we will see that Gaussian measure can be constructed in H , with quite a large group of invariant transformations.

Problems 1.5

23) Let $A = (a_{ij})_{i,j=1}^n$ and $B = (b_{ij})_{i,j=1}^n$ be positive definite matrices. Prove that the matrix $C = (a_{ij}b_{ij})_{i,j=1}^n$ is positive definite as well.

24) Let f and g be pdfs, with cumulative distribution functions F and G , respectively. Prove the following:

a) For each $\alpha \in (-1, 1)$,

$$h(x, y) := f(x)g(y) + \alpha f(x)(1 - 2F(x))g(y)(1 - 2G(y)), \quad (x, y) \in \mathbb{R}^2$$

is pdf, with marginal densities $f(x)$ and $g(y)$.

b) Assume additionally that f and g are even functions and $\int_{\mathbb{R}} |x|f(x)dx < \infty$, $\int_{\mathbb{R}} |y|f(y)dy < \infty$. Let $(X; Y)^T$ be a random vector with pdf equal to $h(x, y)$. If $\alpha \in (0, 1)$, then X and Y are positively correlated, and if $\alpha \in (-1, 0)$, then X and Y are negatively correlated.

25) Based on problem (24), construct Gaussian random variables X and Y which are not jointly Gaussian.

26) Let $A \in \mathbb{R}^{n \times n}$ be a symmetric matrix and $X \sim N(0, S)$ in \mathbb{R}^n . Denote eigenvalues of $S^{1/2}AS^{1/2}$ as $\lambda_1, \dots, \lambda_n$. Prove that $I_\alpha := \mathbb{E} \exp\{\alpha(AX, X)\} < \infty$ if, and only if, $\alpha\lambda_k < \frac{1}{2}$, $k = 1, \dots, n$. Show that in this case $I_\alpha = \frac{1}{\sqrt{\prod_{k=1}^n (1 - 2\alpha\lambda_k)}}$.

27) Let $X \sim N(0, I_n)$. Find for which real α

$$I_\alpha := \mathbb{E} \|X\|^{-\alpha} < \infty.$$